

Statistical estimation of the intermittency coefficient of a random cascade

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Fractals and Related Fields II

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The setting

Estimation procedures and convergence rates

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Statistical inference for multifractal processes

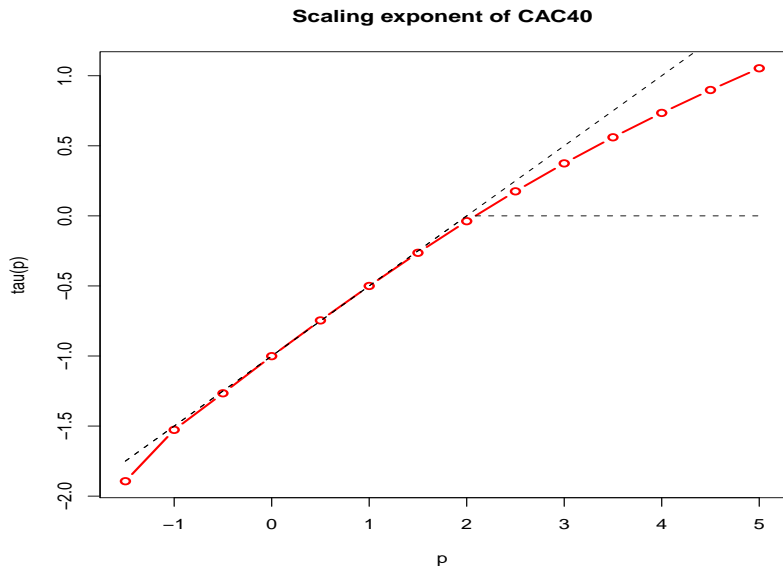
- ▶ Let $X = (X(t), t \geq 0)$ a real-valued random process s.t.

$$\mathbb{E}[|X(t+s) - X(t)|^p] \approx s^{\zeta(p)} \quad \text{as } s \rightarrow 0.$$

- ▶ Hölder inequality: $p \mapsto \zeta(p)$ is a **concave function**. We say that X is **monofractal** if ζ is **linear**, and X is **multifractal** if ζ is **strictly concave**.
→ multifractal formalism, Hölder regularity of the sample paths of X .
- ▶ Suppose that we have some data $X_0, X_{1/n}, \dots, X_1$. Can we recognize if X is mono- or multifractal? **and how accurately can we reconstruct ζ as the number of data n grows ?**

→ **Applications to real data:** turbulence, finance...

Scaling exponent on financial data



Pointwise estimation vs. cumulant estimation

Estimating $\zeta(p)$ for all p in some real interval:

- ▶ First possibility: find consistent estimates of $\zeta(p_1), \dots, \zeta(p_{u_n})$ with $u_n \rightarrow +\infty$ as the number of data n grows.
- ▶ Second possibility: estimate $\zeta(0), \zeta'(0), \zeta''(0), \dots$
Intermittency coefficient: $\zeta''(0)$ which indicates whether ζ is **linear** or not.
- ▶ Here I consider processes of **random cascades** that satisfy a **scaling property**: for some interval I and all $r \in (0, 1)$

$$(X(rt), t \in I) \stackrel{d}{=} r W_r (X(t), t \in I)$$

with W_r a positive r.v. indep. of X . Then $\zeta^{(k)}(0)$ can be recovered from $\mathbb{E}[\log^k W_r]$.

Mandelbrot cascades [Mandelbrot 1974, Kahane and Peyrière 1976]

A **Mandelbrot cascade** is a **continuous, nonnegative and increasing** process $(X(t), t \in [0, T])$. Let W a positive r.v. with $\mathbb{E}[W] = 1$, and

$$(W_i, i \in \{0, 1\}^k, k \in \mathbb{N})$$

i.i.d. copies of W . $X(t)$ is defined as the limit of the sequence

$$X_1(t) = \int_0^t (W_0 \mathbf{1}_{u \in [0, 1/2]} + W_1 \mathbf{1}_{u \in [1/2, 1]}) du$$

$$X_2(t) = \int_0^t (W_0 W_{00} \mathbf{1}_{u \in [0, 1/4]} + W_0 W_{01} \mathbf{1}_{u \in [1/4, 1/2]} \\ + W_1 W_{10} \mathbf{1}_{u \in [1/2, 3/4]} + W_1 W_{11} \mathbf{1}_{u \in [3/4, 1]}) du$$

$$X_3(t) = \dots$$

Log-normal cascades

- ▶ More elaborate framework: **grid free** constructions of random cascades: Kahane (1985), Barral and Mandelbrot (2002), Bacry and Muzy (2003)...
- ▶ Popular setting: "**log-normal**" cascades: $\log W$ is a Gaussian random variable with variance λ^2 .
- ▶ They are multifractal processes with scaling exponent

$$\zeta(p) = p - \lambda^2 p(p-1)/2.$$

- ▶ Hence estimating the **function** ζ is the same as simply estimating the **scalar** λ^2 – the *intermittency coefficient*.

The setting

Estimation procedures and convergence rates

Estimation based on the empirical moments

Idea: approximate the true theoretical moments by empirical moments.

► Let

$$S_n(p) = \frac{1}{n} \sum_{k=0}^{n-1} |X((k+1)/n) - X(k/n)|^p.$$

Then from the multifractal property

$$S_n(p) \approx \mathbb{E}[|X((k+1)/n) - X(k/n)|^p] = n^{-1+\lambda^2 p(p-1)/2}.$$

► Hence define

$$\tilde{\lambda}_n^{2,p} = \frac{2}{p(p-1)} \left(1 + \frac{\log S_n(p)}{\log n} \right)$$

or

$$\hat{\lambda}_n^{2,p} = \frac{2}{p(p-1)} \left(1 + \frac{\log S_{2n}(p) - \log S_n(p)}{\log 2} \right).$$

Convergence rate

Theorem 1 (Ossiander and Waymire 2000, D. 2011, Ludeña and Soulier 2011)

Let X a log-normal Mandelbrot cascade, or a log-normal multifractal random measure (Bacry and Muzy 2003). Then for some $p^ > 1$ and all $p \in (0, p^*)$, $p \neq 1$,*

$$|\hat{\lambda}_n^{2,p} - \lambda^2| \asymp n^{-1/2 + \lambda^2 p^2 / 2}.$$

- ▶ **Slower rate of convergence** than the **usual parametric rate $1/\sqrt{n}$** .

Are there estimation procedures that converge faster ?

Logarithms of increments

Abry, Jaffard, Roux and Wendt (2007); Bacry, Kozhemyak and Muzy (2008) → estimates based on the properties of the **log** of the increments (or wavelet coef., or wavelete leaders.)

- ▶ From the scaling property:

$$(X(t/n), t \in I) \stackrel{d}{=} 1/n W_{1/n} (X(t), t \in I)$$

with $\log W_{1/n}$ a **Gaussian r.v.** with variance $\lambda^2 \log n$.

- ▶ Let $x_{n,k} = \log |X((k+1)/n) - X(k/n)|$, then

$$x_{n,k} \stackrel{d}{=} -\log(n) + w_{n,k} + m_{n,k}, \quad 0 \leq k \leq n-1,$$

with $(w_{n,k})_k$ a **Gaussian sequence** with covariance $\text{Cov}[w_{n,k}, w_{n,k'}] \approx c + \lambda^2 \log(n/(|k-k'|+1))$, while $(m_{n,k})$ can be considered as a "noise".

Using the moments of the log of the increments for estimating λ^2

From the previous decomposition, we hope to recover λ^2 from the second order properties of $w_{n,k}$.

- ▶ Abry *et al.* 2007: estimator based on the **empirical variance** of the $x_{n,k}$'s.
→ I find a rate of convergence $\frac{\log n}{\sqrt{n}}$.
- ▶ Bacry *et al.* 2008: estimator based on the **empirical covariance** of the $x_{N,k}$'s.
→ rate of convergence $\frac{\sqrt{\log n}}{\sqrt{n}}$.
- ▶ D., 2011: estimator based on the **empirical variance** of $X_{N,k+1} - X_{N,k}$.
→ rate of convergence $\frac{1}{\sqrt{n}}$.

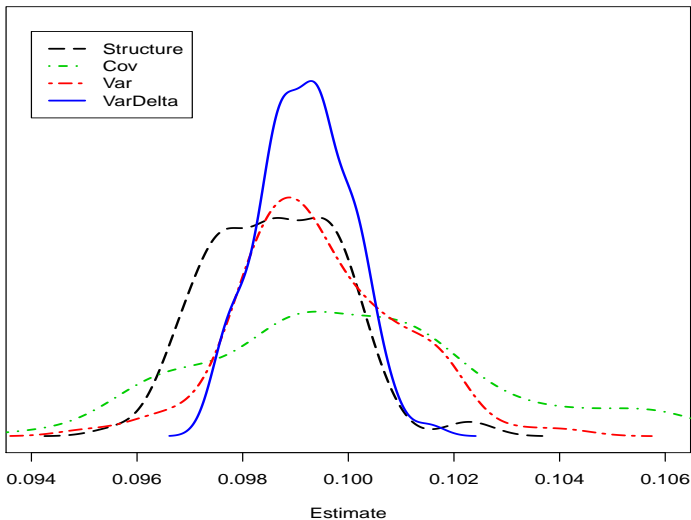


Figure: Empirical distribution of the estimators for 100 simulations of a log-normal MRM; $\lambda^2 = 0.1$; $N = 32\,768$.

Conclusion

- ▶ How accurately can we characterize the multifractality of some data when the number of data grows?
- ▶ In the "simple" case of log-normal random cascades, multifractality is characterized by a single real number, the intermittency coefficient.
- ▶ We find different convergence rates for different estimation procedures of this coefficient.
- ▶ In particular, the "usual" approach based on the empirical moments of the increments of the process is sub-optimal.